

# Statistical Inference in a Hierarchical Growth Curve Model

Takashi Kanda

*Department of Environment Design, Hiroshima Institute of Technology*

*2-1-1, Miyake, Saeki-ku*

*Hiroshima, Japan, 731-5193*

*kanda@cc.it-hiroshma.ac.jp*

We consider a growth curve model with hierarchical types of within-individuals design matrices. This model is an extension of the usual growth curve model introduced by Potthoff and Roy (1964). From a practical view, we state here the extended growth curve model with two hierarchical structures. The model is the following: we observe an  $N \times p$  random matrix  $Y$  whose  $N$  rows are independently distributed, each with a  $p$ -variate normal distribution having unknown positive definite covariance matrix  $\Sigma$ . The means of the elements of  $Y$  are assumed to be of the form

$$E(Y) = A_1 \Xi_1 X_1 + A_2 \Xi_2 X_{(2)} = [A_1 \ A_2] \begin{bmatrix} \Xi_1 & O \\ \Xi_{21} & \Xi_{22} \end{bmatrix} \begin{bmatrix} X_1 \\ X_2 \end{bmatrix},$$

where  $A_i$  are known  $N \times k_i$  design matrices between-individual design matrices of ranks  $k_i$  ( $i=1,2$ ),  $X_1$  and  $X_{(2)} = [X_1' \ X_2']'$  are  $q_1 \times p$  and  $q \times p$  ( $q = q_1 + q_2 \leq p$ ) within-individuals design matrices with rank  $q_1$  and  $q$ , respectively. The parameters  $\Xi_1 = [\Xi_{11}]$  and  $\Xi_2 = [\Xi_{21} \ \Xi_{22}]$  are  $k_1 \times q_1$  and  $k_2 \times q$  unknown matrices, respectively.

For this model, the following problems have been studied by several authors:

- (1) Maximum likelihood estimators and their basic properties
- (2) The derivation of the AIC criterion for the model selection
- (3) Testing the significance of hierarchical structures
- (4) Likelihood ratio test for the hypothesis  $\Xi_{22} = O$
- (5) Simultaneous confidence regions of  $\Xi_1, \Xi_2$  based simplified estimators and maximum likelihood estimators

In this paper, we consider LR test for the hypothesis

$$C \begin{bmatrix} \Xi_{11} & O \\ \Xi_{21} & \Xi_{22} \end{bmatrix} D = O$$

for some  $c \times k$  matrix  $C$  and  $s \times q$  matrix  $D$  under the assumption  $A_1' A_2 = O$ . In particular we restrict ourselves in the following two cases: (i)  $C = [C_1 \ O]$  and  $D = [D_1' \ O]'$ , then the above hypothesis is  $C_1 \Xi_{11} D_1 = O$ , and (ii)  $C = [O \ C_2]$  and  $D = [D_1' \ D_2']'$ , this case is  $C_2 [\Xi_{21} \ \Xi_{22}] D = O$ . Moreover, in the second case we consider the case where  $D$  is the following matrix

$$D = \begin{bmatrix} \mathbf{I}_{q_1} & \mathbf{O} & \mathbf{O} \\ \mathbf{O} & \mathbf{I}_{q_3} & \mathbf{O} \end{bmatrix},$$

where  $q_3 \leq q_2$ . Considering appropriate orthogonal matrices  $H$  and  $B$ , we can transform  $Y$  to a canonical form  $Z = H'YB$  such that  $Z \sim N_{N,p}(E(Z), \Omega \otimes I_N)$  with

$$(a) E(Z) = \begin{bmatrix} \Theta_{11} & \Theta_{12} & \mathbf{O} & \mathbf{O} \\ \Theta_{21} & \Theta_{22} & \mathbf{O} & \mathbf{O} \\ \Theta_{31} & \Theta_{32} & \Theta_{33} & \mathbf{O} \\ \mathbf{O} & \mathbf{O} & \mathbf{O} & \mathbf{O} \end{bmatrix}, \quad (b) E(Z) = \begin{bmatrix} \Theta_{11} & \mathbf{O} & \mathbf{O} \\ \Theta_{21} & \Theta_{22} & \mathbf{O} \\ \Theta_{31} & \Theta_{32} & \mathbf{O} \\ \mathbf{O} & \mathbf{O} & \mathbf{O} \end{bmatrix}$$

and  $\Omega = B\Sigma B'$ , where  $H = [H_1 H_2 H_3 H_4] \in O(N)$  and  $B = [B_1' B_2' B_3' B_4']' \in O(p) : H_i (N \times r_i), B_i (q_i \times p) : r_1 = c_1, r_2 = k_1 + k_2 - c_1, r_3 = N - r_1 - r_2$ . Then the null hypotheses for testing problem (i) and (ii) are equivalent to

$$(i)' \quad \Theta_{11} = \mathbf{O}, \quad (ii)' \quad \Theta_{2(12)} = [\Theta_{21}, \Theta_{22}] = \mathbf{O},$$

respectively. The likelihood ratio statistic for (i)' is given by

$$(c) \quad \Lambda = |S_e| / |S_e + S_h|$$

where  $W = [Z_{41} Z_{42} Z_{43} Z_{44}]' [Z_{41} Z_{42} Z_{43} Z_{44}]$ ,  $S_e = W_{22 \cdot 34} = W_{22} - W_{2(34)} W_{(34)(34)}^{-1} W_{(34)2}$ ,  $S_h = (Z_{12} - Z_{1(34)} W_{(34)(34)}^{-1} W_{(34)2})' (I + Z_{1(34)} W_{(34)(34)}^{-1} Z_{(34)1})^{-1} (Z_{12} - Z_{1(34)} W_{(34)(34)}^{-1} W_{(34)2})$ . Furthermore, the null distribution of this  $\Lambda$  is the ordinary  $\Lambda$ -distribution. Note that  $\Lambda$  does not depend  $Z_2 = [Z_{21} Z_{22} Z_{23} Z_{24}]$  and  $Z_3 = [Z_{31} Z_{32} Z_{33} Z_{34}]$ .

The likelihood ratio statistic for testing problem (ii)' is given by

$$(d) \quad \Lambda = \Lambda^{(1)} \Lambda^{(2)}$$

where  $\Lambda^{(i)} = |S_e^{(i)}| / |S_e^{(i)} + S_h^{(i)}|$ . The explicit expressions of  $S_e^{(i)}$  and  $S_h^{(i)}$  are omitted here. We note that  $\Lambda^{(1)}$  and  $\Lambda^{(2)}$  are the significant test statistics for  $\Theta_{21}$  and  $\Theta_{22}$ , respectively.

Moreover, under the hypothesis,  $\Lambda^{(1)}$  and  $\Lambda^{(2)}$  are mutually independent and are distributed as  $\Lambda$ -distribution.

## REFERENCES

Potthoff, R. F. and Roy, S. N. (1964). A generalized multivariate analysis of variance model useful especially for growth curve problems. *Biometrika* 51, 313 - 326.

## FRENCH RÉSUMÉ

*Dans cet article, il s'agit d'un modèle élargi à courbe de croissance avec des matrices-plan à l'intérieur des individus ayant deux hiérarchies. Nous déduisons une ratio de vraisemblance statistique pour l'hypothèse linéaire de paramètres.*