

# On record values of discrete distributions

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## Abstract

Let  $X_1, X_2, \dots$  be a sequence of independent and identically distributed random variables taking on values  $0, 1, \dots$  with distribution function  $F$  such that  $F(n) < 1$  for any  $n$ , and let  $EX_1 \ln(1 + X_1) < \infty$ . Let  $X_{U(n)}$  be the  $n$ -th record value. In this paper we show that  $E(X_{U(n+2)} - X_{U(n)} | X_{U(n)} = l) = c$ , for some  $n > 0$  and all  $l \geq n - 1$  if and only if  $X_1$  has the distribution of the form

$$p_k = P(X_1 = k) = (1 - p_0) \left(1 - \frac{2}{c}\right) \left(\frac{2}{c}\right)^{k-1}, k = 1, 2, 3, \dots$$

( $p_0 = P(X_1 = 0)$  is arbitrary  $0 < p_0 < 1$ ). Our result is generalization of the theorems Srivastava (1979), Korwar (1984) and Aliev (2000).

**Key Words:** records, characterization of discrete distributions, characterization of geometric distribution.

## 1. INTRODUCTION

A lot of papers in the field of records are devoted to characterizations of distributions via records (see references below and references in the book Ahsanullah (1995) and in the paper Nevzorov (1987) among many others). Great interest in records exists because we often come across with them in our everyday life in such a way that singling out and fixing of record values.

Let  $X_1, X_2, \dots$  be a sequence of independent and identically distributed random variables (r.v.'s) taking on values  $0, 1, \dots$  with distribution function  $F$  such that  $F(n) < 1$  for any  $n$ , and let  $EX_1 \ln(1 + X_1) < \infty$ . Define the sequence of weak record times  $L(n)$ , record times  $U(n)$ , weak record values  $X_{L(n)}$  and record values  $X_{U(n)}$  as follows:

$$\begin{aligned} L(1) &= 1, & L(n+1) &= \min\{j > L(n) : X_j \geq X_{L(n)}\}, & n &= 1, 2, \dots, \\ U(1) &= 1, & U(n+1) &= \min\{j > U(n) : X_j > X_{U(n)}\}, & n &= 1, 2, \dots \end{aligned}$$

The following are well-known:

1)  $X_1$  has a distribution of the form  $P\{X_1 \geq m\} = \prod_{i=1}^m \frac{\alpha+(i-1)\beta}{1+\alpha+i\beta}$  for some  $\alpha > 0, \beta \geq 0$ , and  $m = 1, 2, \dots$ , if and only if  $E(X_{L(n+1)} - X_{L(n)} | X_{L(n)} = l) = \alpha + \beta l$ , for some  $n \geq 1$  and for all  $l = 0, 1, \dots$  (Stepanov (1994)). If  $\beta = 0$  this result corresponds to the geometric distribution;

2) if  $\{S_i\}_{i=0}^{\infty}$  is any sequence of positive numbers such that  $S_i > S_{i-1} - 1$  for all  $i$  and  $\prod_{i=1}^{\infty} \frac{S_i}{1+S_i} = 0$  then  $X_1$  has distribution of the form  $P(X_1 \geq m) = \prod_{i=1}^m \frac{S_{i-1}}{1+S_i}$  for all  $m = 1, 2, \dots$  if and only if  $E(X_{L(n+1)} - X_{L(n)} | X_{L(n)} = l) = S_l$  for some  $n \geq 1$  and for all  $l = 0, 1, \dots$  (Aliev (1998)). In the case of  $S_l = \alpha + \beta l$  from this result implies the above result of Stepanov (1994);

3) if  $\{S_i\}_{i=0}^{\infty}$  is any sequence of positive finite numbers such that  $S_i > S_{i-1} - 1$  for all  $i$  and if there exists  $F(x)$  such that  $E(X_{L(n+2)} - X_{L(n)} | X_{L(n)} = l) = S_l$  for some  $n \geq 1$  and all  $l = 0, 1, \dots$  then  $F(x)$  is unique (Aliev (1999)). There are some examples of characterization theorems for known distributions in this paper and it is easy to extend these characterizations for other distributions. Note also that the characterization problem using  $E\{X_{L(n+k)} - X_{L(n)} | X_{L(n)} = l\}$  for  $k > 2$  is still open;

4)  $X_1$  has a distribution of the form  $p_m = (1-p_0) \left(\frac{\alpha-1}{\alpha}\right)^{m-1} \frac{1}{\alpha}$  for some  $0 < p_0 < 1$  and  $m = 1, 2, \dots$ , if and only if  $E(X_{U(n+1)} - X_{U(n)} | X_{U(n)} = l) = \alpha$ , for some  $n \geq 1, \alpha > 1$  and for all  $l = 0, 1, \dots$  (Srivastava (1979));

5)  $X_1$  has a distribution of the form  $p_1 = (1-p_0) \frac{1+\beta}{\alpha+\beta}$ ,  $p_k = (1-p_0) \frac{1+\beta}{\alpha+k\beta} \prod_{j=1}^{k-1} \left(1 - \frac{1+\beta}{\alpha+j\beta}\right)$  for some  $0 < p_0 < 1$  and  $m = 2, 3, \dots$ , if and only if  $E(X_{U(n+1)} - X_{U(n)} | X_{U(n)} = l) = \alpha + \beta l$ , for some  $n \geq 1, \alpha > 1, \beta \geq 0$  and for all  $l = 0, 1, \dots$  (Korwar (1984));

6) if  $\{S_i\}_{i=0}^{\infty}$  is any sequence of positive numbers such that  $S_k > 1, S_{k+1} > S_k - 1$  for all  $k$  and  $\prod_{i=0}^{\infty} \frac{S_i-1}{S_i} = 0$  then  $X_1$  has distribution of the form  $p_1 = (1-p_0) \left(1 - \frac{S_0-1}{S_1}\right)$ ,  $p_k = (1-p_0) \prod_{j=1}^{k-1} \frac{S_{j-1}-1}{S_j} \left(1 - \frac{S_{k-1}-1}{S_k}\right)$ ,  $k = 2, 3, \dots$  for arbitrary  $0 < p_0 < 1$  ( $p_j = P(X_1 = j)$ ) if and only if  $E(X_{U(n+1)} - X_{U(n)} | X_{U(n)} = l) = S_l$ , for some  $n > 0$  and all  $l \geq n - 1$  (Aliev (2000)). This result is generalization of the theorems Srivastava (1979) and Korwar (1984).

Similar results for continuous distributions are obtained in the papers of Nagaraja (1977), Blaquez and Rebollo (1997), Ahsanullah and Wesolowski (1998) and Demblinka and Wesolowski (1999).

## 2. MAIN RESULT

**Theorem.** Let  $X_{U(n)}$  be the  $n$ -th record value and  $c > 0$ . A necessary and suffi-

cient condition for a r.v.  $X_1$  to have distribution of the form

$$p_k = P(X_1 = k) = (1 - p_0) \left(1 - \frac{2}{c}\right) \left(\frac{2}{c}\right)^{k-1}, k = 1, 2, 3, \dots$$

$$(p_0 = P(X_1 = 0) \text{ is arbitrary } 0 < p_0 < 1)$$

is

$$E(X_{U(n+2)} - X_{U(n)} | X_{U(n)} = l) = c, \text{ for some } n > 0 \text{ and all } l \geq n - 1.$$

**Remark 1.** The characterization problems when the regression  $E\{X_{U(n+k)} - X_{U(n)} | X_{U(n)} = s\}$  is not constant for  $k = 2$  and the case  $k > 2$  are steel open.

**Remark 2.** The proof of the theorem is based on the technique of the paper Aliev (2000) and may be extended for the case  $k = 2$ , when  $E\{X_{U(n+k)} - X_{U(n)} | X_{U(n)} = s\}$  is not constant.

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